Examiner: Xiangfeng Yang (013-285788). Things allowed: a calculator, an English-Swedish dictionary. Scores rating (Betygsgränser): 8-11 points giving rate 3; 11.5-14.5 points giving rate 4; 15-18 points giving rate 5.

1 (3 points)

Let X and Y be independent Exp(1)-distributed random variables. Show that X/(X+Y) and X+Y are independent. and find their density functions $f_{X/(X+Y)}(u)$ and $f_{X+Y}(v)$.

2 (3 points)

Let Y be a Binomial random variable with a random parameter X as follows:

 $Y \mid X = x \sim Bin(n, x),$ with $X \sim U(0, 1).$ [This can be also written as $Y \mid X \sim Bin(n, X)$]

Compute the expectation E(Y) of Y and the covariance Cov(X, Y) of X and Y.

(3 points)3

Prove, with the aid of suitable transforms, that if $X \sim Bin(n, p)$ and $Y \sim Bin(m, p)$ are independent, then $X + Y \sim Bin(n+m, p).$

(3 points)4

The random variables $X_1, X_2, \ldots, X_n, Y_1, Y_2, \ldots, Y_n$ are independent U(0, 1)-distributed.

(4.1) Find the density function $f_{X_{(n)}}(x)$ of $X_{(n)}$, where $X_{(n)} = \max\{X_1, X_2, \dots, X_n\}$. (4.2) Find the probability $P\left(\frac{\max\{X_{(n)}, Y_{(n)}\}}{\min\{X_{(n)}, Y_{(n)}\}} \ge 2\right)$.

$\mathbf{5}$ (3 points)

Let $\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \sim N(\mathbf{0}, \mathbf{\Lambda})$ with $\mathbf{\Lambda} = \begin{pmatrix} 3 & 1 \\ 1 & 2 \end{pmatrix}$, that is, the density function is $f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{2\pi \cdot \sqrt{\det(\mathbf{\Lambda})}} \exp\left\{-\frac{1}{2}\mathbf{x}'\mathbf{\Lambda}^{-1}\mathbf{x}\right\}$. Find the conditional density function $f_{X_1+X_2|X_1-X_2=0}(x)$ of $X_1 + X_2$ given that $X_1 - X_2 = 0$. (Hints: you might need to use the inverse formula $\begin{pmatrix} a & b \\ c & d \end{pmatrix}^{-1} = \frac{1}{ad-bc} \begin{pmatrix} d & -b \\ -c & a \end{pmatrix}$.)

(3 points)6

(6.1) Let $P(Z_n = 0) = 1 - \frac{1}{n}$, $P(Z_n = 1) = \frac{1}{2n}$ and $P(Z_n = -1) = \frac{1}{2n}$ for $n \ge 1$. Prove that $Z_n \xrightarrow{p} 0$. (6.2) Let Y be a Normal random variable with a random parameter X as follows:

 $Y \mid X = x \sim N(0, x), \quad \text{ with } X \sim Po(\lambda). \quad \Big[\text{This can be also written as } Y \mid X \sim N(0, X) \Big]$

Prove that $Y/\sqrt{\lambda} \xrightarrow{d} N(0,1)$ as $\lambda \to \infty$.

(Hint: use the characteristic function of $Y/\sqrt{\lambda}$, and a fact $e^x = 1 + x + o(x)$ as $x \to 0$.)

Followingis a list of discrete distribu An asterisk (*) indicates that the e	ttions, abbreviations, their probability functions, i expression is too complicated to present here; in s	means, va some case	ariances, and es a closed fo	l characteristic functio ormula does not even	ons. exist.
Distribution, notation	Probability function	E X	$\operatorname{Var} X$	$\varphi_X(t)$	
One point $\delta(a)$	p(a) = 1	в	0	e^{ita}	
Symmetric Bernoulli	$p(-1) = p(1) = \frac{1}{2}$	0	1	$\cos t$	
Bernoulli $\operatorname{Be}(p), 0 \leq p \leq 1$	$p(0) = q, \ p(1) = p; \ q = 1 - p$	d	bd	$q + pe^{it}$	
Binomial Bin $(n, p), n = 1, 2, \dots, 0 \le p \le 1$	$p(k) = {n \choose k} p^k q^{n-k}, \ k = 0, 1, \dots, n; \ q = 1 - p$	du	bdu	$(q + pe^{it})^n$	
Geometric $\operatorname{Ge}(p), \ 0 \leq p \leq 1$	$p(k) = pq^k, \ k = 0, 1, 2, \dots; \ q = 1 - p$	$\frac{d}{d}$	$\frac{q}{p^2}$	$\frac{p}{1-qe^{it}}$	
First success $\operatorname{Fs}(p), 0 \leq p \leq 1$	$p(k) = pq^{k-1}, \ k = 1, 2, \dots; \ q = 1 - p$	$\frac{1}{p}$	$p^{\frac{q}{2}}$	$\frac{pe^{it}}{1-qe^{it}}$	
Negative binomial NBin $(n, p), n = 1, 2, 3, \dots, 0 \le p \le 1$	$p(k) = {n+k-1 \choose k} p^n q^k, \ k = 0, 1, 2, \dots;$ q = 1 - p	$\frac{d}{b}u$	$n \frac{q}{p^2}$	$\big(\frac{p}{1-q^{e^{it}}}\big)^n$	
Poisson $Po(m), m > 0$	$p(k) = e^{-m} \; rac{m^k}{k!}, \; k = 0, 1, 2, \ldots$	m	m	$e^{m(e^{it}-1)}$	
Hypergeometric $H(N, n, p), n = 0, 1, \dots, N,$ $N = 1, \frac{2}{N}, \dots, 1$ $p = 0, \frac{1}{N}, \frac{2}{N}, \dots, 1$	$p(k) = \frac{\binom{Np}{k}\binom{Nq}{n-k}}{\binom{N}{n}}, k = 0, 1, \dots, Np;$ $q = 1 - p;$ $n - k = 0, \dots, Nq$	du	$npq \frac{N-n}{N-1}$	*	

Discrete Distributions

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An asterisk (*) indicate	s that the expression is too complicated to j	present here	; in some cases a close	d formula does not even
Distribution, notation	Density	E X	$\operatorname{Var} X$	$\varphi_X(t)$
Uniform/Rectangular U(a, b)	$f(x) = \frac{1}{b-a}, \ a < x < b$	$\frac{1}{2}(a+b)$	$\frac{1}{12}(b-a)^2$	$\frac{e^{itb} - e^{ita}}{it(b-a)}$
U(0,1) U(-1,1)	$f(x) = 1, \ 0 < x < 1$ $f(x) = \frac{1}{2}, \ x < 1$	- <mark>1</mark> -	3 <mark>1- 12</mark>	$\frac{e^{it}-1}{it}$
Triangular Tri (a,b)	$f(x) = \frac{2}{b-a} \left(1 - \frac{2}{b-a} \left x - \frac{a+b}{2} \right \right)$ a < x < b	$\frac{1}{2}(a+b)$	$\frac{1}{24}(b-a)^2$	$\left(\frac{e^{itb/2}-e^{ita/2}}{\frac{1}{2}it(b-a)}\right)^2$
$\operatorname{Tri}(-1,1)$	$f(x) = 1 - x , \ x < 1$	0	- I 0	$\left(\frac{\sin\frac{t}{2}}{\frac{t}{2}}\right)^2$
Exponential $Exp(a), a > 0$	$f(x) = \frac{1}{a} e^{-x/a}, \ x > 0$	a	a^2	$\frac{1}{1-ait}$
Gamma $\Gamma(p,a), \ a > 0, \ p > 0$	$f(x) = rac{1}{\Gamma(p)} x^{p-1} rac{1}{a^p} e^{-x/a}, \; x > 0$	ра	pa^2	$\frac{1}{(1-ait)^p}$
Chi-square $\chi^2(n), n = 1, 2, 3, \dots$	$f(x) = \frac{1}{\Gamma(\frac{n}{2})} x^{\frac{1}{2}n-1} \left(\frac{1}{2}\right)^{n/2} e^{-x/2}, \ x > 0$	u	2n	$\frac{1}{(1-2it)^{n/2}}$
Laplace $L(a), a > 0$	$f(x)=rac{1}{2a}e^{- x /a}, \ -\infty < x < \infty$	0	$2a^2$	$\frac{1}{1+a^2t^2}$
Beta $\beta(r,s), r,s > 0$	$f(x) = \frac{\Gamma(r+s)}{\Gamma(r)\Gamma(s)} x^{r-1} (1-x)^{s-1},$	$\frac{r}{r+s}$	$\frac{rs}{(r+s)^2(r+s+1)}$	*
	0 < x < 1			

nces, and characteristic functions. abbreviations their densities "" distributions. list of s Following is a

Continuous Distributions

Distribution, notation	Density	E X	$\operatorname{Var} X$	$\varphi_X(t)$
Weibull $W(lpha,eta), lpha,eta>0$	$f(x) = rac{1}{lpha eta} x^{(1/eta) - 1} e^{-x^{1/eta} / lpha}, \; x > 0$	$lpha^eta\Gamma(eta+1)$	$a^{2eta}ig(\Gamma(2eta+1)\ -\Gamma(eta+1)^2ig)$	*
Rayleigh Ra $(\alpha), \alpha > 0$	$f(x) = \frac{2}{\alpha} x e^{-x^2/\alpha}, \ x > 0$	$\frac{1}{2}\sqrt{\pi\alpha}$	$lpha(1-rac{1}{4}\pi)$	*
Normal $\begin{split} & N(\mu,\sigma^2), \\ & -\infty < \mu < \infty, \sigma > 0 \end{split}$	$f(x)=rac{1}{\sigma\sqrt{2\pi}}e^{-rac{1}{2}(x-\mu)^2/\sigma^2},$	Ц	σ^2	$e^{i\mu t-rac{1}{2}t^2\sigma^2}$
	$-\infty < x < \infty$			
N(0,1)	$f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, -\infty < x < \infty$	0	Ι	$e^{-t^{2}/2}$
Log-normal $LN(\mu, \sigma^2), -\infty < \mu < \infty, \ \sigma > 0$	$f(x) = \frac{1}{\sigma x \sqrt{2\pi}} e^{-\frac{1}{2}(\log x - \mu)^2 / \sigma^2}, \ x > 0$	$e^{\mu+rac{1}{2}\sigma^2}$	$e^{2\mu} \left(e^{2\sigma^2} - e^{\sigma^2} ight)$	*
(Student's) t $t(n), n = 1, 2, \dots$	$f(x) = rac{\Gamma(rac{n+1}{2})}{\sqrt{\pi n} \Gamma(rac{n}{2})} \cdot drac{1}{(1+rac{n-1}{2})^{(n+1)/2}}, \ -\infty < x < \infty$	0	$\frac{n}{n-2},n>2$	*
(Fisher's) F $F(m \ n) \ m \ n = 1$ 2	$f(x) = \frac{\Gamma(\frac{m+n}{2})(\frac{m}{n})^{m/2}}{\Gamma(\frac{m}{2})\Gamma(\frac{n}{2})} \cdot \frac{x^{m/2-1}}{(1+\frac{mx}{n})^{(m+n)/2}},$	$rac{n}{n-2},$	$rac{n^2(m+2)}{m(n-2)(n-4)} - \left(rac{n}{n-2} ight)^2,$	*
···· (= (+	x > 0	n > 2	n > 4	

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Continuous Distributions (continued)

B Some Distributions and Their Characteristics

Distribution, notation	Density	E X	$\operatorname{Var} X$	$\varphi_X(t)$
Cauchy				
C(m,a)	$f(x) = \frac{1}{\pi} \cdot \frac{a}{a^2 + (x-m)^2}, \ -\infty < x < \infty$	Ŕ	Ā	$e^{imt-a t }$
C(0,1)	$f(x) = \frac{1}{\pi} \cdot \frac{1}{1+x^2}, -\infty < x < \infty$	R	R	$e^{- t }$
Pareto	$f(x)=rac{lpha k^lpha}{x^{lpha+1}},\ x>k$	$\frac{\alpha k}{\alpha - 1}, \alpha > 1$	$\frac{\alpha k^2}{(\alpha-2)(\alpha-1)^2}, \alpha > 2,$	*
$\operatorname{Pa}(k,\alpha), k > 0, \alpha > 0$				

Continuous Distributions (continued)