Examiner: Xiangfeng Yang (013-285788). Things allowed: a calculator, an English-Swedish dictionary.
Scores rating (Betygsgränser): 8-11 points giving rate 3; 11.5-14.5 points giving rate 4; 15-18 points giving rate 5 .

## 1 (3 points)

Suppose that $X$ and $Y$ have a joint probability density function as follows

$$
f(x, y)= \begin{cases}\frac{1}{y} \cdot e^{-x / y} \cdot e^{-y}, & \text { for } 0<x<\infty \text { and } 0<y<\infty \\ 0, & \text { otherwise }\end{cases}
$$

Are $X / Y$ and $Y$ independent? Why?

## 2 (3 points)

Let $X$ be a Binomial random variable with a random parameter $N$ as follows:

$$
X \mid N=n \sim \operatorname{Bin}(n, p), \quad \text { with } N \sim \operatorname{Po}(\lambda)
$$

Find the probability $P(X=k)$ for $k=0,1,2, \ldots$

## 3 (3 points)

Let $X$ be a discrete random variable with

$$
P(X=0)=0.2, \quad P(X=1)=0.5, \quad P(X=2)=0.3 .
$$

(3.1) (1p) Find the probability generating function $g_{X}(t)=E\left(t^{X}\right)$.
(3.2) (1p) Find the moment generating function $\psi_{X}(t)=E\left(e^{t X}\right)$.
(3.3) (1p) Find the characteristic generating function $\varphi_{X}(t)=E\left(e^{i t X}\right)$.

## 4 (3 points)

Let $X_{1} \sim \operatorname{Exp}(1)$ and $X_{2} \sim \operatorname{Exp}(1)$ be two independent exponential random variables. Set $X_{(1)}=\min \left\{X_{1}, X_{2}\right\}$ and $X_{(2)}=\max \left\{X_{1}, X_{2}\right\}$. Find the conditional expectation $E\left(X_{(2)} \mid X_{(1)}=x\right)$.

## 5 (3 points)

Let $X_{1}, X_{2}, \ldots$ be independent and identically distributed random variables with mean 0 and variance 1 . Assume that $N \sim \operatorname{Po}(\lambda)$ is independent of $X_{1}, X_{2}, \ldots$. Show that

$$
\frac{X_{1}+X_{2}+\ldots+X_{N}}{\sqrt{N}} \xrightarrow{d} N(0,1) \quad \text { as } \lambda \rightarrow \infty .
$$

## 6 (3 points)

Let $X_{1}, X_{2}, \ldots$ be independent $U(0,1)$ random variables, and set $Y_{n}=X_{1} \cdot X_{2} \cdot \ldots \cdot X_{n}$ for $n=1,2, \ldots$
(6.1) (1p) Prove that $Y_{n} \xrightarrow{p} 0$ as $n \rightarrow \infty$.
(6.2) (2p) Prove that $Y_{n} \xrightarrow{\text { a.s. }} 0$ as $n \rightarrow \infty$.
Discrete Distributions
Followingis a list of discrete distributions, abbreviations, their probability functions, means, variances, and characteristic functions. An asterisk $\left({ }^{*}\right)$ indicates that the expression is too complicated to present here; in some cases a closed formula does not even exist.
Probability function $E X \quad \operatorname{Var} X \quad \varphi_{X}(t)$



$0 \quad$
1
$p q$
$n p q$
0



$0 \quad 0$
0
$p$
ह
01 $2 \quad-$ $\qquad$
-12 $p(a)=1$ $p(-1)=p(1)=\frac{1}{2}$ O
$p(k)=\binom{n}{k} p^{k} q^{n-k}, k=0,1, \ldots, n ; q=1-p$ $p(k)=p q^{k}, k=0,1,2, \ldots ; q=1-p$
 ,$\ldots$
$k=0,1, \ldots, N p ;$ $k=0,1, \ldots$,
$\quad q=1-p ;$
$n-k=0, \ldots, N q$


Continuous Distributions

Continuous Distributions (continued)

| Distribution, notation | Density | $E X$ | $\operatorname{Var} X$ | $\varphi_{X}(t)$ |
| :---: | :---: | :---: | :---: | :---: |
| Weibull $W(\alpha, \beta), \alpha, \beta>0$ | $f(x)=\frac{1}{\alpha \beta} x^{(1 / \beta)-1} e^{-x^{1 / \beta} / \alpha}, x>0$ | $\alpha^{\beta} \Gamma(\beta+1)$ | $\begin{aligned} & a^{2 \beta}(\Gamma(2 \beta+1) \\ & \left.\quad-\Gamma(\beta+1)^{2}\right) \end{aligned}$ | * |
| Rayleigh $\operatorname{Ra}(\alpha), \alpha>0$ | $f(x)=\frac{2}{\alpha} x e^{-x^{2} / \alpha}, x>0$ | $\frac{1}{2} \sqrt{\pi \alpha}$ | $\alpha\left(1-\frac{1}{4} \pi\right)$ | * |
| Normal $\begin{aligned} & N\left(\mu, \sigma^{2}\right) \\ & -\infty<\mu<\infty, \sigma>0 \end{aligned}$ | $f(x)=\frac{1}{\sigma \sqrt{2 \pi}} e^{-\frac{1}{2}(x-\mu)^{2} / \sigma^{2}},$ $-\infty<x<\infty$ | $\mu$ | $\sigma^{2}$ | $e^{i \mu t-\frac{1}{2} t^{2} \sigma^{2}}$ |
| $N(0,1)$ | $f(x)=\frac{1}{\sqrt{2 \pi}} e^{-x^{2} / 2},-\infty<x<\infty$ | 0 | 1 | $e^{-t^{2} / 2}$ |
| Log-normal $\begin{aligned} & L N\left(\mu, \sigma^{2}\right), \\ & -\infty<\mu<\infty, \sigma>0 \end{aligned}$ | $f(x)=\frac{1}{\sigma x \sqrt{2 \pi}} e^{-\frac{1}{2}(\log x-\mu)^{2} / \sigma^{2}}, x>0$ | $e^{\mu+\frac{1}{2} \sigma^{2}}$ | $e^{2 \mu}\left(e^{2 \sigma^{2}}-e^{\sigma^{2}}\right)$ | * |
| (Student's) $t$ $t(n), n=1,2, \ldots$ | $f(x)=\frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{\pi n \Gamma\left(\frac{n}{2}\right)}} \cdot d \frac{1}{\left(1+\frac{x^{2}}{n}\right)^{(n+1) / 2}},$ $-\infty<x<\infty$ | 0 | $\frac{n}{n-2}, n>2$ | * |
| $\begin{aligned} & \text { (Fisher's) } F \\ & \quad F(m, n), m, n=1,2, \end{aligned}$ | $f(x)=\frac{\Gamma\left(\frac{m+n}{2}\right)\left(\frac{m}{n}\right)^{m / 2}}{\Gamma\left(\frac{m}{2}\right) \Gamma\left(\frac{n}{2}\right)} \cdot \frac{x^{m / 2-1}}{\left(1+\frac{m x}{n}\right)^{(m+n) / 2}},$ $x>0$ | $\begin{aligned} & \frac{n}{n-2}, \\ & n>2 \end{aligned}$ | $\begin{array}{r} \frac{n^{2}(m+2)}{m(n-2)(n-4)}-\left(\frac{n}{n-2}\right)^{2}, \\ n>4 \end{array}$ | * |

Continuous Distributions (continued)

| Distribution, notation | Density | $E X$ | $\operatorname{Var} X$ | $\varphi_{X}(t)$ |
| :--- | :--- | :---: | :---: | :---: |
| Cauchy |  |  |  |  |
| $\quad C(m, a)$ | $f(x)=\frac{1}{\pi} \cdot \frac{a}{a^{2}+(x-m)^{2}},-\infty<x<\infty$ | $\nexists$ | $A$ | $e^{i m t-a\|t\|}$ |
| $\quad C(0,1)$ | $f(x)=\frac{1}{\pi} \cdot \frac{1}{1+x^{2}},-\infty<x<\infty$ | $A$ | $A$ | $e^{-\|t\|}$ |
| Pareto | $f(x)=\frac{\alpha k^{\alpha}}{x^{\alpha+1}}, x>k$ | $\frac{\alpha k}{\alpha-1}, \alpha>1$ | $\frac{\alpha k^{2}}{(\alpha-2)(\alpha-1)^{2}}, \alpha>2$, | $*$ |
| $\operatorname{Pa}(k, \alpha), k>0, \alpha>0$ |  |  |  |  |

