

Examiner: Xiangfeng Yang (Tel: 013 28 57 88).**Things allowed (Hjälpmedel):** a calculator.**Scores rating (Betygsgränser):** 8-11 points giving rate 3; 11.5-14.5 points giving rate 4; 15-18 points giving rate 5.

1 (3 points)

Two events A and B are independent with $P(A) = 0.3$ and $P(B) = 0.5$.

- (1.1) (0.5p) Find the probability $P(A \cap B)$.
- (1.2) (0.5p) Find the probability $P(A \cup B)$.
- (1.3) (1p) Find the conditional probability $P(A | A \cup B)$.
- (1.4) (1p) Find the probability $P(A' \cap B)$, where A' is the complement of A .

Solution. (1.1) It follows from independence that $P(A \cap B) = P(A) \cdot P(B) = 0.3 \cdot 0.5 = 0.15$.

(1.2) With the result in (1.1), it holds that

$$P(A \cup B) = P(A) + P(B) - P(A \cap B) = 0.3 + 0.5 - 0.15 = 0.65.$$

(1.3) The conditional probability is computed as follows (with the result in (1.2)):

$$P(A | A \cup B) = \frac{P(A \cap (A \cup B))}{P(A \cup B)} = \frac{P(A)}{P(A \cup B)} = \frac{0.3}{0.65} = 0.4615.$$

(1.4) Since A and B are independent, A' and B are also independent. Therefore

$$P(A' \cap B) = P(A') \cdot P(B) = (1 - P(A)) \cdot P(B) = (1 - 0.3) \cdot 0.5 = 0.35.$$

□

2 (3 points)

A PHEV (plug-in hybrid electric vehicle) car has two engines: a diesel engine and an electric engine. Let X be lifespan (in years) of the diesel engine, and Y lifespan (in years) of the electric engine. Assume that $X \sim \text{Exp}(\frac{1}{10})$ and $Y \sim \text{Exp}(\frac{1}{12})$ are independent and exponential random variables, namely X has mean 10 years and Y has mean 12 years.

(2.1) (1.5p) Find the probability that the electric engine fails before the diesel engine, namely, $P(X > Y)$.

(2.2) (1.5p) Find the probability that the total lifespan of the PHEV is at least 25 years, namely, $P(X + Y \geq 25)$.

Solution. The probability density function of X is $f_X(x) = \frac{1}{10}e^{-x/10}$ for $x \geq 0$, and the probability density function of Y is $f_Y(y) = \frac{1}{12}e^{-y/12}$ for $y \geq 0$. Since X and Y are independent, the joint probability density function of (X, Y) is

$$f(x, y) = f_X(x) \cdot f_Y(y) = \frac{1}{10}e^{-x/10} \cdot \frac{1}{12}e^{-y/12}, \quad \text{for } x \geq 0, y \geq 0.$$

(2.1)

$$\begin{aligned} P(X > Y) &= \int \int_{\{x>y\}} f(x, y) dx dy = \int_0^\infty \left(\int_0^x f(x, y) dy \right) dx \\ &= \int_0^\infty \frac{1}{10}e^{-x/10} \left(\int_0^x \frac{1}{12}e^{-y/12} dy \right) dx = \int_0^\infty \frac{1}{10}e^{-x/10} \left(1 - e^{-x/12} \right) dx \\ &= \int_0^\infty \frac{1}{10}e^{-x/10} dx - \int_0^\infty \frac{1}{10}e^{-x(1/10+1/12)} dx = 1 - \frac{\frac{1}{10}}{1/10 + 1/12} = \frac{10}{22} = \frac{5}{11} = 0.4545. \end{aligned}$$

(2.2)

$$\begin{aligned}
P(X + Y \geq 25) &= \int \int_{\{x+y \geq 25\}} f(x, y) dx dy = 1 - \int \int_{\{x+y < 25\}} f(x, y) dx dy \\
&= 1 - \int_0^{25} \left(\int_0^{25-x} f(x, y) dy \right) dx \\
&= 1 - \int_0^{25} \frac{1}{10} e^{-x/10} \left(\int_0^{25-x} \frac{1}{12} e^{-y/12} dy \right) dx \\
&= 1 - \int_0^{25} \frac{1}{10} e^{-x/10} \left(1 - e^{-(25-x)/12} \right) dx \\
&= 1 - \int_0^{25} \frac{1}{10} e^{-x/10} dx + \int_0^{25} \frac{1}{10} e^{-x/10 - (25-x)/12} dx \\
&= 1 - \int_0^{25} \frac{1}{10} e^{-x/10} dx + \int_0^{25} \frac{1}{10} e^{-25/12} e^{-x(1/10 - 1/12)} dx \\
&= 1 - (1 - e^{-25/10}) + \frac{1}{10} e^{-25/12} \int_0^{25} e^{-x/60} dx \\
&= e^{-25/10} + \frac{1}{10} e^{-25/12} (60 - 60e^{-25/60}) \\
&= e^{-2.5} + 6e^{-2.0833} - 6e^{-2.5} = 6e^{-2.0833} - 5e^{-2.5} = 6 \cdot 0.1245 - 5 \cdot 0.0821 = 0.747 - 0.4105 = 0.3365.
\end{aligned}$$

□

3 (3 points)

Every week Emma goes to gym a few times. Let X be the distribution of numbers of times that Emma goes to gym in one week. Assume that X is a discrete random variable with the following probability mass function:

X	1	2	3
$p(x)$	0.2	0.5	0.3

For example, $P(X = 1) = 0.2$ means that in one week Emma goes to gym only 1 time with probability 20%.

(3.1) (1p) Find the mean $\mu = E(X)$ and variance $\sigma^2 = V(X)$ of X .

(3.2) (2p) Assume that gym activities for Emma in different weeks are independent. Find the probability that Emma goes to gym at least 120 times in the next 52 weeks? (Hint: Central Limit Theorem (CLT))

Solution. (3.1)

$$\begin{aligned}
\mu &= E(X) = 1 \cdot 0.2 + 2 \cdot 0.5 + 3 \cdot 0.3 = 2.1, \\
\sigma^2 &= V(X) = E(X^2) - \mu^2 = 1^2 \cdot 0.2 + 2^2 \cdot 0.5 + 3^2 \cdot 0.3 - 2.1^2 = 4.9 - 4.41 = 0.49.
\end{aligned}$$

(3.2) Let X_1, X_2, \dots, X_{52} denote the times Emma goes to gym in the next 52 weeks, it is then from CLT that

$$\begin{aligned}
P(\text{Emma goes to gym at least 120 times in the next 52 weeks}) &= P(X_1 + X_2 + \dots + X_{52} \geq 120) \\
&= P(\bar{X} \geq 120/52) = P\left(\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \geq \frac{2.31 - 2.1}{\sigma/\sqrt{n}}\right) \\
&= P(N(0, 1) \geq \frac{2.31 - 2.1}{\sqrt{0.49}/\sqrt{52}}) = P(N(0, 1) \geq 2.16) \\
&= 1 - P(N(0, 1) < 2.16) = 1 - \Phi(2.16) = 1 - 0.9846 = 0.0154.
\end{aligned}$$

□

4 (3 points)

(4.1) (1p) A population X is a continuous random variable with a probability density function

$$f(x) = \frac{1}{a-10}, \quad \text{for } 10 \leq x \leq a,$$

where $a > 10$ is unknown. A sample $\{x_1, x_2, \dots, x_n\}$ is taken from this population. Use the method of moments to find a point estimate \hat{a}_{MM} of a .

(4.2) (2p) Another population Y is a continuous random variable with a probability density function

$$f(y) = \frac{1}{b-c}, \quad \text{for } c \leq y \leq b,$$

where b and c (with $b > c$) are unknown. A sample $\{y_1, y_2, \dots, y_n\}$ is taken from this population. Use the maximum-likelihood method to find point estimates \hat{b}_{ML} and \hat{c}_{ML} of b and c .

Solution. (4.1) The mean $E(X)$ of the population can be computed as

$$E(X) = \int_{10}^a x f(x) dx = \int_{10}^a \frac{x}{a-10} dx = \frac{a+10}{2}.$$

It is then from the equation $E(X) = \bar{x}$ that

$$\frac{a+10}{2} = \bar{x} \Rightarrow \hat{a}_{MM} = 2\bar{x} - 10.$$

(4.2) The likelihood function is

$$L(b, c) = f(y_1) \cdot f(y_2) \cdot \dots \cdot f(y_n) = \left(\frac{1}{b-c} \right)^n = (b-c)^{-n}, \quad \text{for } c \leq y_i \leq b, 1 \leq i \leq n.$$

Then non-trivial domain “ $c \leq y_i \leq b, 1 \leq i \leq n$ ” can be rewritten as

$$\begin{aligned} c &\leq \min\{y_1, y_2, \dots, y_n\}, \text{ that is, } c \in (-\infty, \min\{y_1, y_2, \dots, y_n\}], \\ b &\geq \max\{y_1, y_2, \dots, y_n\}, \text{ that is, } b \in [\max\{y_1, y_2, \dots, y_n\}, +\infty). \end{aligned}$$

- The partial derivative $\frac{\partial L(b,c)}{\partial b} = -n(b-c)^{-n-1} < 0$, so $L(b, c)$ is decreasing in b . Therefore

$$\hat{b}_{ML} = \max\{y_1, y_2, \dots, y_n\}.$$

- The partial derivative $\frac{\partial L(b,c)}{\partial c} = n(b-c)^{-n-1} > 0$, so $L(b, c)$ is increasing in c . Therefore

$$\hat{c}_{ML} = \min\{y_1, y_2, \dots, y_n\}.$$

□

5 (3 points)

A new medicine is to be introduced. It is believed that the risk of side effects is affected by the dose size. To study this, 600 patients are randomly selected and divided into two groups with each group consisting of 300 patients. Two different dose sizes are given to these two groups and the results are:

Groups	Dose sizes	How many patients have side effects?
Group A: 300 patients	0.20	54
Group B: 300 patients	0.30	95

Let p_A and p_B denote the proportions of patients having side effects with the above mentioned two dose sizes. Does the above sample provide any evidence that the risk of side effects is affected by the dose size? Answer this using an appropriate two-sided 95% confidence interval. (Hint: Do NOT use the method of Hypotheses Test!).

Solution. The two sample proportions are

$$\hat{p}_A = \frac{54}{300}, \quad \hat{p}_B = \frac{95}{300}.$$

The 95% confidence interval of $p_A - p_B$ is (with $n_1 = n_2 = 300$)

$$\begin{aligned} I_{p_A - p_B} &= (\hat{p}_A - \hat{p}_B) \mp z_{\alpha/2} \sqrt{\frac{\hat{p}_A(1 - \hat{p}_A)}{n_1} + \frac{\hat{p}_B(1 - \hat{p}_B)}{n_2}} \\ &= \left(\frac{54}{300} - \frac{95}{300} \right) \mp 1.96 \sqrt{\frac{\frac{54}{300}(1 - \frac{54}{300})}{300} + \frac{\frac{95}{300}(1 - \frac{95}{300})}{300}} \\ &= -0.1367 \mp 1.96 \cdot 0.0348 = -0.1367 \mp 0.0682 = (-0.2049, -0.0685). \end{aligned}$$

Since both -0.2049 and -0.0685 are negative, it suggests that $p_A - p_B < 0$ which means $p_A < p_B$. (One may say that higher dose increases the risk of side effects). □

6 (3 points)

A coffee supplier has plenty of 500 g coffee boxes. The coffee supplier claims that the distribution X of weights of all these coffee boxes is $X \sim N(500, 2^2)$. To check this claim, a randomly selected 100 coffee boxes give the following:

Weight of coffee boxes	< 498	498 – 502	> 502
Number of coffee boxes	39	45	16

The above results say that, among the randomly selected 100 coffee boxes: there are 39 coffee boxes whose weights are less than 498 g; there are 45 coffee boxes whose weights are between 498 g and 502 g, and there are 16 coffee boxes whose weights are more than 502 g. Use χ^2 -test with a significance level 5% to check whether or not the sample provides any evidence showing the coffee supplier's claim is wrong.

Solution. The hypotheses to be checked are:

$$H_0 : X \sim N(500, 2^2) \quad \text{against} \quad H_1 : X \not\sim N(500, 2^2).$$

To this end, under H_0 , the probabilities for these three groups can be computed as follows:

$$\begin{aligned} p_1 &= P(X < 498) = P(N(500, 2^2) < 498) = P\left(\frac{N(500, 2^2) - 500}{2} < \frac{498 - 500}{2}\right) \\ &\quad = P(N(0, 1) < -1) = \Phi(-1) = 1 - \Phi(1) = 1 - 0.8413 = 0.1587. \\ p_2 &= P(498 \leq X < 502) = P(498 \leq N(500, 2^2) \leq 502) = P\left(\frac{498 - 500}{2} \leq N(0, 1) \leq \frac{502 - 500}{2}\right) \\ &\quad = P(-1 \leq N(0, 1) \leq 1) = \Phi(1) - \Phi(-1) = \Phi(1) - (1 - \Phi(1)) = 2\Phi(1) - 1 = 0.6826. \\ p_3 &= P(X > 502) = P(N(500, 2^2) > 502) = P\left(\frac{N(500, 2^2) - 500}{2} > \frac{502 - 500}{2}\right) \\ &\quad = P(N(0, 1) > 1) = 1 - P(N(0, 1) \leq 1) = 1 - \Phi(1) = 1 - 0.8413 = 0.1587. \end{aligned}$$

The TS can be now evaluated as

$$\begin{aligned} TS &= \sum_{i=1}^3 \frac{(n_i - np_i)^2}{np_i} = \frac{(39 - 100 \cdot 0.1587)^2}{100 \cdot 0.1587} + \frac{(45 - 100 \cdot 0.6826)^2}{100 \cdot 0.6826} + \frac{(16 - 100 \cdot 0.1587)^2}{100 \cdot 0.1587} \\ &= 33.711 + 7.926 + 0.001 = 41.638. \end{aligned}$$

The rejection region is

$$C = (\chi^2_{\alpha}(k-1), \infty) = (5.99, \infty).$$

Since $TS \in C$, we reject H_0 . Namely, the sample provides evidence that the coffee supplier's claim is wrong ($X \not\sim N(500, 2^2)$). □

1. Basic probability

(1.1) Conditional probability $P(A|B) = \frac{P(A \cap B)}{P(B)}$.

(1.2) Total probability $P(B) = \sum_{i=1}^k P(B|A_i)P(A_i)$ where $\{A_i\}$ are disjoint and $\cup_{i=1}^k A_i = S$.

(1.3) Bayes' Theorem $P(A_j|B) = \frac{P(B|A_j)P(A_j)}{\sum_{i=1}^k P(B|A_i)P(A_i)}$ where $\{A_i\}$ are in (1.2).

2. Random variables (r.v.s)

(2.1) Discrete r.v. X has a pmf $p(x) = P(X = x)$ satisfying $p(x) \geq 0$ and $\sum p(x_i) = 1$,

X	x_1	x_2	\cdots	x_n	\cdots
$p(x)$	$p(x_1)$	$p(x_2)$	\cdots	$p(x_n)$	\cdots

Expectation (or *Expected value* or *mean*) $\mu_X = E(X) = \sum x_i p(x_i)$;
 Variance $\sigma_X^2 = V(X) = E(X - \mu_X)^2 = E(X^2) - \mu_X^2 = \sum x_i^2 p(x_i) - (\sum x_i p(x_i))^2$.

(2.2) Continuous r.v. X has a pdf $f(x)$ satisfying $f(x) \geq 0$ and $\int_{-\infty}^{\infty} f(x)dx = 1$,

$$P(a < X < b) = \int_a^b f(x)dx.$$

Expectation (or *Expected value* or *mean*) $\mu_X = E(X) = \int_{-\infty}^{\infty} xf(x)dx$;

Variance $\sigma_X^2 = V(X) = E(X - \mu_X)^2 = E(X^2) - \mu_X^2 = \int_{-\infty}^{\infty} x^2 f(x)dx - (\int_{-\infty}^{\infty} xf(x)dx)^2$.

(2.3) Cumulative distribution function (cdf) of a r.v. X is $F(x) = P(X \leq x)$.

(2.4) X and Y are r.v.s, a , b and c are scalars, then

$$E(aX + bY + c) = aE(X) + bE(Y) + c,$$

$$V(aX + bY + c) = a^2 V(X) + b^2 V(Y) + 2ab \text{cov}(X, Y),$$

$$E(g(X, Y)) = \begin{cases} \sum_{i,j} g(x_i, y_j) \cdot p(x_i, y_j), & \text{for discrete } (X, Y), \\ \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) \cdot f(x, y)dxdy, & \text{for continuous } (X, Y). \end{cases}$$

(2.5) • Discrete r.v. (X, Y) has a joint pmf $p(x, y)$ satisfying $p(x, y) \geq 0$ and $\sum_{x_i} \sum_{y_i} p(x_i, y_i) = 1$.

The marginal pmf of X is $p_X(x) = \sum_y p(x, y)$;
 The marginal pmf of Y is $p_Y(y) = \sum_x p(x, y)$;

X and Y are *independent* if $p(x, y) = p_X(x) \cdot p_Y(y)$.

• Continuous r.v. (X, Y) has a joint pdf $f(x, y)$ satisfying $f(x, y) \geq 0$ and $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y)dxdy = 1$.

The marginal pdf of X is $f_X(x) = \int_{-\infty}^{\infty} f(x, y)dy$;

The marginal pdf of Y is $f_Y(y) = \int_{-\infty}^{\infty} f(x, y)dx$;
 X and Y are *independent* if $f(x, y) = f_X(x) \cdot f_Y(y)$.

3. Several special r.v.s

(3.1) $X \sim Bin(n, p)$ has a pmf $p(x) = P(X = x) = \binom{n}{x} \cdot p^x \cdot (1-p)^{n-x}$, $x = 0, 1, 2, \dots, n$.

$$E(X) = n \cdot p, \quad V(X) = n \cdot p \cdot (1-p).$$

(3.2) $X \sim Po(\lambda)$ has a pmf $p(x) = P(X = x) = \frac{e^{-\lambda} \lambda^x}{x!}$, $x = 0, 1, 2, \dots$.

$$E(X) = \lambda, \quad V(X) = \lambda.$$

(3.3) $X \sim Hypergeometric$ has a pmf $p(x) = P(X = x) = \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}$.

(3.4) $X \sim Exp(\lambda)$ has a pdf

$$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0, \\ 0, & \text{otherwise.} \end{cases}$$

(3.5) $X \sim N(\mu, \sigma^2)$ has a pdf

$$f(x) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, -\infty < x < \infty.$$

$$\begin{aligned} E(X) &= \mu, & V(X) &= \sigma^2 \\ (3.6) \quad X &\sim U(a, b) \text{ has a pdf} & \end{aligned}$$

$$f(x) = \begin{cases} \frac{1}{b-a}, & a < x < b, \\ 0, & \text{otherwise.} \end{cases}$$

$$E(X) = \frac{a+b}{2}, \quad V(X) = \frac{(b-a)^2}{12}.$$

4. Central Limit Theorem (CLT)

Suppose that a population has mean = μ and variance = σ^2 . A random sample $\{X_1, X_2, \dots, X_n\}$ from this population is given. Then for large $n \geq 30$,

$$\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1).$$

- If the population is normal, then (1) holds for any n .
- Note that $\mu = E(\bar{X})$ and $(\sigma/\sqrt{n})^2 = V(\bar{X})$.

5. Several notations in statistics

(5.1) Sample mean: $\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n} = \frac{\sum X_i}{n}$; $\bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n} = \frac{\sum x_i}{n}$.

(5.2) Sample variance:

$$S^2 = \frac{\sum (X_i - \bar{X})^2}{n-1} = \frac{1}{n-1} \left(\sum X_i^2 - \frac{(\sum X_i)^2}{n} \right); \quad s^2 = \frac{\sum (x_i - \bar{x})^2}{n-1} = \frac{1}{n-1} \left(\sum x_i^2 - \frac{(\sum x_i)^2}{n} \right).$$

• Capital letters \bar{X} and S^2 refer to the objects based on random sample (therefore they are in general r.v.s), while small letters \bar{x} and s^2 are the objects based on observations (so they are scalars).

(5.3) A point estimator of θ obtained by Maximum Likelihood method is denoted as $\hat{\theta}_{ML}$.

6. Confidence Interval (CI)

In this course, three types of confidence intervals are studied depending on the unknown population parameter(s): CI-1 (confidence intervals for population mean(s)), CI-2 (confidence intervals for population variance(s)), and CI-3 (confidence intervals for population proportion(s)).

CI-1: $(1 - \alpha)$ CI of a population mean μ

case 1.1 (any n) If population $X \sim N(\mu, \sigma^2)$ and σ^2 is known, then $\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1)$ and

$$I_\mu = \left(\bar{x} - z_{\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}, \bar{x} + z_{\alpha/2} \cdot \frac{\sigma}{\sqrt{n}} \right) := \bar{x} \mp z_{\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}.$$

case 1.2 ($n \geq 30$) For any population X , it holds that $\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1)$ and

$$I_\mu = \bar{x} \mp z_{\alpha/2} \cdot \frac{\sigma}{\sqrt{n}} \text{ or } I_\mu = \bar{x} \mp z_{\alpha/2} \cdot \frac{\hat{\sigma}}{\sqrt{n}}.$$

case 1.3 (any n) If population $X \sim N(\mu, \sigma^2)$ and σ^2 is unknown, then $\frac{\bar{X} - \mu}{S/\sqrt{n}} \sim T(n-1)$ and

$$I_\mu = \bar{x} \mp t_{\alpha/2}(n-1) \cdot \frac{s}{\sqrt{n}}.$$

CI-1': $(1 - \alpha)$ CI of the difference of two population means $\mu_X - \mu_Y$

case 1.1' (any n_1, n_2) If independent populations $X \sim N(\mu_X, \sigma_X^2)$, $Y \sim N(\mu_Y, \sigma_Y^2)$, and σ_X^2, σ_Y^2 are known, then $\frac{(\bar{X} - \bar{Y}) - (\mu_X - \mu_Y)}{\sqrt{\frac{\sigma_X^2}{n_1} + \frac{\sigma_Y^2}{n_2}}} \sim N(0, 1)$, and $I_{\mu_X - \mu_Y} = (\bar{x} - \bar{y}) \mp z_{\alpha/2} \cdot \sqrt{\frac{\sigma_X^2}{n_1} + \frac{\sigma_Y^2}{n_2}}$.

case 1.2' ($n_1, n_2 \geq 30$) For any independent populations X and Y , it holds that

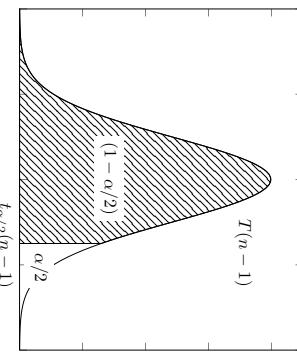
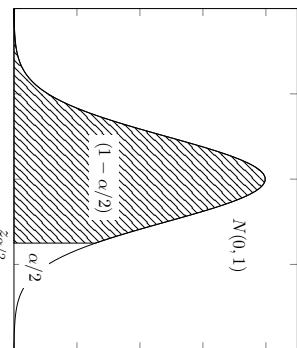
$$\frac{(\bar{X} - \bar{Y}) - (\mu_X - \mu_Y)}{\sqrt{\frac{\sigma_X^2}{n_1} + \frac{\sigma_Y^2}{n_2}}} \sim N(0, 1) \text{ and}$$

$$I_{\mu_X - \mu_Y} = (\bar{x} - \bar{y}) \mp z_{\alpha/2} \cdot \sqrt{\frac{\sigma_X^2}{n_1} + \frac{\sigma_Y^2}{n_2}} \text{ or } I_{\mu_X - \mu_Y} = (\bar{x} - \bar{y}) \mp z_{\alpha/2} \cdot \sqrt{\frac{\hat{\sigma}_X^2}{n_1} + \frac{\hat{\sigma}_Y^2}{n_2}}.$$

case 1.3' (any n_1, n_2) If independent populations $X \sim N(\mu_X, \sigma_X^2)$, $Y \sim N(\mu_Y, \sigma_Y^2)$, where σ_X^2, σ_Y^2 are unknown but $\sigma_X^2 = \sigma_Y^2$, then

$$\frac{(\bar{X} - \bar{Y}) - (\mu_X - \mu_Y)}{\sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim T(n_1 + n_2 - 2), \text{ where } S^2 = \frac{(n_1 - 1)S_X^2 + (n_2 - 1)S_Y^2}{n_1 + n_2 - 2}, \text{ and}$$

$$I_{\mu_X - \mu_Y} = (\bar{x} - \bar{y}) \mp t_{\alpha/2}(n_1 + n_2 - 2) \cdot s \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}.$$



CI-2: $(1 - \alpha)$ CI of population variance(s) σ^2

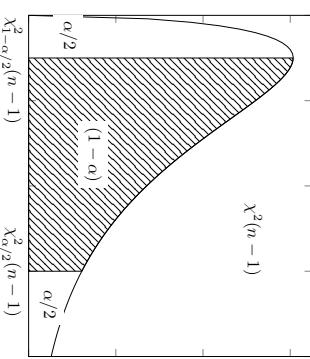
- If a population $X \sim N(\mu, \sigma^2)$ and σ^2 is unknown, then $\frac{(n-1)S^2}{\sigma^2} \sim \chi^2(n-1)$, and

$$I_{\sigma^2} = \left(\frac{(n-1)s^2}{\chi_{\alpha/2}^2(n-1)}, \frac{(n-1)s^2}{\chi_{1-\alpha/2}^2(n-1)} \right).$$

- If two independent populations $X \sim N(\mu_X, \sigma^2)$ and $Y \sim N(\mu_Y, \sigma^2)$, and σ^2 is unknown, then $\frac{(n_1+n_2-2)S^2}{\sigma^2} \sim \chi^2(n_1 + n_2 - 2)$, and

$$I_{\sigma^2} = \left(\frac{(n_1 + n_2 - 2)s^2}{\chi_{\alpha/2}^2(n_1 + n_2 - 2)}, \frac{(n_1 + n_2 - 2)s^2}{\chi_{1-\alpha/2}^2(n_1 + n_2 - 2)} \right),$$

where $S^2 = \frac{(n_1 - 1)S_X^2 + (n_2 - 1)S_Y^2}{n_1 + n_2 - 2}$.



CI-3: $(1 - \alpha)$ CI of population proportion(s)

- If a (large) population has an unknown proportion p , then $\frac{\hat{p} - p}{\sqrt{p(1-p)/n}} \sim N(0, 1)$ if $n\hat{p} \geq 10$ and $n(1 - \hat{p}) \geq 10$ with $\hat{p} = x/n$, and $I_p = \hat{p} \mp z_{\alpha/2} \cdot \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$.
- If two independent (large) populations have unknown proportions p_1 and p_2 , then

$$\frac{(\hat{p}_1 - \hat{p}_2) - (p_1 - p_2)}{\sqrt{\frac{p_1(1-p_1)}{n_1} + \frac{p_2(1-p_2)}{n_2}}} \sim N(0, 1)$$

if $n_i\hat{p}_i \geq 10$ and $n_i(1 - \hat{p}_i) \geq 10$ for $i = 1, 2$, and $I_{p_1 - p_2} = (\hat{p}_1 - \hat{p}_2) \mp z_{\alpha/2} \cdot \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$.

7. Hypothesis Test (HT)

	H_0 is true	H_0 is false and $\theta = \theta_1$
reject H_0	(type I error or significance level) α	(power) $h(\theta_1) = 1 - h(\theta_1)$
don't reject H_0	1 - α	(type II error) $\beta(\theta_1) = 1 - h(\theta_1)$

reject $H_0 \Leftrightarrow TS \in C \Leftrightarrow p\text{-value} < \alpha$

χ^2 tests for populations (non-parametric tests)

Suppose that for a random sample of a population X , the n elements of it are classified into k disjoint groups $A_i, 1 \leq i \leq k$. For each group $A_i, 1 \leq i \leq k$, suppose that there are $N_i, 1 \leq i \leq k$ elements inside. Let $p_i = P(A_i)$ assuming a given distribution of X . Note that $p_1 + p_2 + \dots + p_k = 1$ and $N_1 + N_2 + \dots + N_k = n$. One wants to test the hypotheses

$$H_0 : P(A_i) = p_i, \quad 1 \leq i \leq k, \quad H_a : P(A_i) \neq p_i \text{ for some } 1 \leq i \leq k.$$

If n is large in the sense that $np_i \geq 5$ for all $1 \leq i \leq k$, then the test statistic is

$$\sum_{i=1}^k \frac{(N_i - np_i)^2}{np_i} \approx \chi^2(k-1).$$

Therefore the observation of the test statistic is

$$TS = \sum_{i=1}^k \frac{(n_i - np_i)^2}{np_i}, \text{ where } n_i \text{ is the observation of } N_i, 1 \leq i \leq k.$$

For the critical region C , one can take (note that if H_0 is true, then TS should be close to zero)

$$C = (\chi_\alpha^2(k-1), \infty).$$

The conclusion would be $TS \in C \iff H_0$ is rejected.

8. Linear and logistic regression

(Multiple) linear regression: $Y = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k + \varepsilon, \varepsilon \sim N(0, \sigma^2)$.

\bullet Y : response variable (which is normal r.v.), $\{x_1, \dots, x_k\}$: predictors (which are scalars).

- sample: $\{(x_{11}, \dots, x_{1k}; y_1), (x_{21}, \dots, x_{2k}; y_2), \dots, (x_{n1}, \dots, x_{nk}; y_n)\}$.
- how to estimate $\beta_j \approx \hat{\beta}_j$: least square method, that is, to minimize $\sum_{i=1}^n (\hat{y}_i - y_i)^2$, where the estimated (multiple) linear regression line \hat{y} is

$$\hat{y} = \beta_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_k x_k.$$

- $\frac{\hat{\beta}_j - \beta_j}{s.e(\hat{\beta}_j)} \sim T(n-k-1)$, this helps determine whether or not the real $\beta_j = 0$?

- $\sigma^2 \approx \frac{SSE}{n-k-1}$, this gives an estimation of the size of the error.

$\bullet R^2 = \frac{SSR}{SST}$ this gives how well the model is (if $R^2 \approx 1$, then the model fits the sample very well).

- How to test $\beta_1 = \dots = \beta_k = 0$? Use the random variable $\frac{SSR/k}{SSE/(n-k-1)} \sim F(k, n-k-1)$.

Logistic regression: Let Y can only take 0 or 1 with $P(Y=1) = p$ and $P(Y=0) = 1-p$,

$$E(Y) = p(x_1, \dots, x_k) = \frac{e^{\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k}}{1 + e^{\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k}}.$$

- Y : response variable (which is Bernoulli r.v. $P(Y=1) = p$ and $P(Y=0) = 1-p$, so $E(Y) = p$), $\{x_1, \dots, x_k\}$: predictors (which are scalars).
- sample: $\{(x_{11}, \dots, x_{1k}; y_1), (x_{21}, \dots, x_{2k}; y_2), \dots, (x_{n1}, \dots, x_{nk}; y_n)\}$.
- how to estimate $\beta_j \approx \hat{\beta}_j$: maximal likelihood method (maximize $\prod_{i=1}^n p(x_{i1}, \dots, x_{ik})^{y_i} (1 - p(x_{i1}, \dots, x_{ik}))^{1-y_i}$).
- $\frac{\hat{\beta}_j - \beta_j}{s.e(\hat{\beta}_j)} \approx N(0, 1)$ for large $n \geq 30$, this helps determine whether or not the real $\beta_j = 0$?
- Classification of a new object $Y(x_1, \dots, x_k)$ as 1 or 0 according

$$Y(x_1, \dots, x_k) = \begin{cases} 1, & \text{if } \hat{p}(x_1, \dots, x_k) \geq 0.5, \\ 0, & \text{if } \hat{p}(x_1, \dots, x_k) < 0.5, \end{cases}$$

where the estimated logit function $\hat{p}(x_1, \dots, x_k)$ is

$$\hat{p}(x_1, \dots, x_k) = \frac{e^{\hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_k x_k}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_k x_k}}.$$

9. Tables

(9.1) Table for $N(0,1)$ standard normal random variable $\Phi(x) = P(N(0,1) \leq x)$, $x \geq 0$.

There is an important relation $\Phi(-x) = 1 - \Phi(x)$, $x \geq 0$.

x	0	1	2	3	4	5	6	7	8	9
0.0	0.5000	0.5040	0.5080	0.5120	0.5160	0.5199	0.5239	0.5279	0.5319	0.5359
0.1	0.5398	0.5438	0.5478	0.5517	0.5557	0.5596	0.5636	0.5675	0.5714	0.5753
0.2	0.5793	0.5832	0.5871	0.5910	0.5948	0.5987	0.6026	0.6064	0.6103	0.6141
0.3	0.6179	0.6217	0.6255	0.6293	0.6331	0.6368	0.6406	0.6443	0.6480	0.6517
0.4	0.6554	0.6591	0.6628	0.6664	0.6700	0.6736	0.6772	0.6808	0.6844	0.6879
0.5	0.6915	0.6950	0.6985	0.7019	0.7054	0.7088	0.7123	0.7157	0.7190	0.7224
0.6	0.7257	0.7291	0.7324	0.7357	0.7389	0.7422	0.7454	0.7486	0.7517	0.7549
0.7	0.7580	0.7611	0.7642	0.7673	0.7704	0.7734	0.7764	0.7794	0.7823	0.7852
0.8	0.7881	0.7910	0.7939	0.7957	0.8023	0.8051	0.8078	0.8106	0.8133	0.8160
0.9	0.8159	0.8186	0.8212	0.8238	0.8264	0.8289	0.8315	0.8340	0.8365	0.8389
1.0	0.8413	0.8438	0.8461	0.8485	0.8508	0.8531	0.8554	0.8577	0.8599	0.8621
1.1	0.8643	0.8665	0.8686	0.8708	0.8729	0.8749	0.8770	0.8790	0.8810	0.8830
1.2	0.8849	0.8869	0.8888	0.8907	0.8925	0.8944	0.8962	0.8980	0.8997	0.9015
1.3	0.9032	0.9049	0.9066	0.9082	0.9099	0.9115	0.9131	0.9147	0.9162	0.9177
1.4	0.9192	0.9222	0.9256	0.9251	0.9265	0.9279	0.9292	0.9306	0.9319	0.9330
1.5	0.9332	0.9345	0.9357	0.9370	0.9382	0.9394	0.9406	0.9418	0.9429	0.9441
1.6	0.9452	0.9463	0.9474	0.9484	0.9495	0.9505	0.9515	0.9525	0.9535	0.9545
1.7	0.9554	0.9564	0.9573	0.9582	0.9591	0.9599	0.9608	0.9616	0.9625	0.9633
1.8	0.9641	0.9649	0.9656	0.9664	0.9671	0.9678	0.9686	0.9693	0.9699	0.9706
1.9	0.9713	0.9719	0.9726	0.9732	0.9738	0.9744	0.9750	0.9756	0.9761	0.9767
2.0	0.9772	0.9778	0.9783	0.9788	0.9793	0.9798	0.9803	0.9808	0.9812	0.9817
2.1	0.9821	0.9826	0.9830	0.9834	0.9838	0.9842	0.9846	0.9850	0.9854	0.9857
2.2	0.9861	0.9864	0.9868	0.9871	0.9875	0.9878	0.9881	0.9884	0.9887	0.9890
2.3	0.9893	0.9896	0.9898	0.9901	0.9904	0.9906	0.9909	0.9911	0.9913	0.9916
2.4	0.9918	0.9920	0.9922	0.9925	0.9927	0.9929	0.9931	0.9932	0.9934	0.9936
2.5	0.9938	0.9940	0.9941	0.9943	0.9945	0.9946	0.9948	0.9949	0.9951	0.9952
2.6	0.9955	0.9956	0.9957	0.9959	0.9960	0.9961	0.9962	0.9963	0.9964	0.9965
2.7	0.9965	0.9966	0.9967	0.9968	0.9969	0.9970	0.9971	0.9972	0.9973	0.9974
2.8	0.9974	0.9975	0.9977	0.9978	0.9978	0.9979	0.9979	0.9980	0.9981	0.9981
2.9	0.9981	0.9982	0.9983	0.9984	0.9984	0.9985	0.9985	0.9986	0.9986	0.9986
3.0	0.9987	0.9987	0.9987	0.9988	0.9988	0.9989	0.9989	0.9990	0.9990	0.9990
3.1	0.9990	0.9991	0.9991	0.9991	0.9992	0.9992	0.9993	0.9993	0.9993	0.9993
3.2	0.9993	0.9993	0.9994	0.9994	0.9994	0.9994	0.9995	0.9995	0.9995	0.9995
3.3	0.9995	0.9995	0.9995	0.9996	0.9996	0.9996	0.9996	0.9996	0.9997	0.9997
3.4	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9998	0.9998
3.5	0.9998	0.9998	0.9998	0.9998	0.9998	0.9998	0.9998	0.9998	0.9998	0.9998
3.6	0.9998	0.9998	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999
3.7	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999
3.8	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999
3.9	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
4.0	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000

(9.2) Table for $T(f)$ random variable $F(x) = P(T(f) \leq x)$, where f is a parameter called 'degrees of freedom'.

f	0.75	0.90	0.95	0.975	0.99	0.995	0.9975	0.9995	$F(x)$
1	1.00	3.08	6.31	12.71	31.82	63.66	127.32	636.62	
2	0.82	1.89	2.92	4.30	6.96	9.92	14.09	31.60	
3	0.76	1.64	2.35	3.18	4.54	5.84	7.45	12.92	
4	0.74	1.53	2.13	2.78	3.75	4.60	5.60	8.61	
5	0.73	1.48	2.02	2.57	3.36	4.03	4.77	6.87	
6	0.72	1.44	2.45	3.14	3.71	4.32	5.96	10.47	
7	0.71	1.41	1.89	2.36	3.00	3.50	4.03	5.41	
8	0.71	1.40	1.86	2.31	2.90	3.36	3.83	5.04	
9	0.70	1.38	1.83	2.26	2.82	3.25	3.69	4.78	
10	0.70	1.37	1.81	2.23	2.76	3.17	3.58	4.59	
11	0.70	1.36	1.80	2.20	2.72	3.11	3.50	4.44	
12	0.70	1.36	1.78	2.18	2.68	3.05	3.43	4.32	
13	0.69	1.35	1.77	2.16	2.65	3.01	3.37	4.22	
14	0.69	1.35	1.76	2.14	2.62	2.98	3.33	4.14	
15	0.69	1.34	1.75	2.13	2.60	2.95	3.29	4.07	
16	0.69	1.34	1.75	2.12	2.58	2.92	3.25	4.01	
17	0.69	1.33	1.74	2.11	2.57	2.90	3.22	3.97	
18	0.69	1.33	1.73	2.10	2.55	2.88	3.20	3.92	
19	0.69	1.33	1.73	2.09	2.54	2.86	3.17	3.88	
20	0.69	1.33	1.72	2.09	2.53	2.85	3.15	3.85	
21	0.69	1.32	1.72	2.08	2.52	2.83	3.14	3.82	
22	0.69	1.32	1.72	2.07	2.51	2.82	3.12	3.79	
23	0.69	1.32	1.71	2.07	2.50	2.81	3.10	3.77	
24	0.68	1.32	1.71	2.06	2.49	2.80	3.09	3.75	
25	0.68	1.32	1.71	2.06	2.49	2.79	3.08	3.73	
26	0.68	1.31	1.71	2.06	2.48	2.78	3.07	3.71	
27	0.68	1.31	1.70	2.05	2.47	2.77	3.06	3.69	
28	0.68	1.31	1.70	2.05	2.47	2.76	3.05	3.67	
29	0.68	1.31	1.70	2.05	2.46	2.76	3.04	3.66	
30	0.68	1.31	1.70	2.04	2.46	2.75	3.03	3.65	
40	0.68	1.30	1.68	2.02	2.42	2.70	2.97	3.55	
50	0.68	1.30	1.68	2.01	2.40	2.68	2.94	3.50	
60	0.68	1.30	1.67	2.00	2.39	2.66	2.91	3.46	
100	0.68	1.29	1.66	2.00	2.36	2.63	2.87	3.39	
∞	0.67	1.28	1.65	1.96	2.33	2.58	2.81	3.29	

(9.3) Table for $\chi^2(f)$ random variable $F(x) = P(\chi^2(f) \leq x)$, where f is a parameter.

f	$F(x)$										
	0.005	0.001	0.005	0.01	0.025	0.05	0.10	0.20	0.30	0.40	0.50
1	0.00	0.00	0.00	0.00	0.00	0.02	0.06	0.15	0.27	0.45	1
2	0.00	0.00	0.01	0.02	0.05	0.10	0.21	0.45	0.71	1.39	2
3	0.02	0.02	0.07	0.11	0.22	0.35	0.58	1.01	1.42	2.37	3
4	0.06	0.09	0.21	0.30	0.48	0.71	1.06	1.65	2.19	2.75	4
5	0.16	0.21	0.41	0.55	0.83	1.15	1.61	2.34	3.00	3.66	5
6	0.30	0.38	0.68	0.87	1.24	1.64	2.20	3.07	3.83	4.57	6
7	0.48	0.60	0.99	1.24	1.69	2.17	2.83	3.82	4.67	5.49	7
8	0.71	0.86	1.34	1.65	2.18	2.73	3.49	4.59	5.53	6.42	8
9	0.97	1.15	1.73	2.09	2.70	3.33	4.17	5.38	6.39	7.36	9
10	1.26	1.48	2.16	2.56	3.25	3.94	4.87	6.18	7.27	8.30	10
11	1.59	1.83	2.60	3.05	3.82	4.57	5.58	6.99	8.15	9.24	11
12	1.93	2.21	3.07	3.57	4.40	5.23	6.30	7.81	9.03	10.18	12
13	2.31	2.62	3.57	4.11	5.01	5.89	7.04	8.63	9.93	11.13	13
14	2.70	3.04	4.07	4.66	5.63	6.57	7.79	9.47	10.82	12.08	14
15	3.11	3.48	4.60	5.23	6.26	7.26	8.55	10.31	11.72	13.03	15
16	3.54	3.94	5.14	5.81	6.91	7.96	9.31	11.15	12.62	13.98	16
17	3.98	4.42	5.70	6.41	7.56	8.67	10.09	12.00	13.53	14.94	17
18	4.44	4.90	6.26	7.01	8.23	9.39	10.86	12.86	14.44	15.89	18
19	4.91	5.41	6.84	7.63	8.91	10.12	11.65	13.72	15.35	16.85	19
20	5.40	5.92	7.43	8.26	9.59	10.85	12.44	14.58	16.27	17.81	20
21	5.90	6.45	8.03	8.90	10.28	11.59	13.24	15.44	17.18	18.77	21
22	6.40	6.98	8.64	9.54	10.98	12.34	14.04	16.31	18.10	19.73	22
23	6.92	7.53	9.26	10.20	11.69	13.09	14.85	17.19	19.02	20.69	23
24	7.45	8.08	9.89	10.86	12.40	13.85	15.66	18.06	19.94	21.65	24
25	7.99	8.65	10.52	11.52	13.12	14.61	16.47	18.94	20.87	22.62	25
26	8.54	9.22	11.16	12.20	13.84	15.38	17.29	19.82	21.79	23.58	26
27	9.09	9.80	11.81	12.88	14.57	16.15	18.11	20.70	22.72	24.54	27
28	9.66	10.39	12.46	13.56	15.31	16.93	18.94	21.59	23.65	25.51	28
29	10.23	10.99	13.12	14.26	16.05	17.71	19.77	22.48	24.58	26.48	29
30	10.80	11.59	13.79	14.95	16.79	18.49	20.60	23.36	25.51	27.44	30
40	16.91	17.92	20.71	22.16	24.43	26.51	29.05	32.34	34.87	37.13	40
50	23.46	24.67	27.99	29.71	32.36	34.76	37.69	41.45	44.31	46.86	50
60	30.34	31.74	35.53	37.48	40.48	43.19	46.46	50.64	53.81	56.62	60
100	59.90	61.92	67.33	70.06	74.22	77.93	82.36	87.95	92.13	95.81	100

Table for $\chi^2(f)$ random variable $F(x) = P(\chi^2(f) \leq x)$, where f is a parameter.

f	$F(x)$									
	0.60	0.70	0.80	0.90	0.95	0.975	0.99	0.995	0.999	0.9995
1	0.71	1.07	1.64	2.71	3.84	5.02	6.63	7.88	10.83	12.12
2	1.83	2.41	3.22	4.61	5.99	7.38	9.21	10.60	13.82	15.20
3	2.95	3.66	4.64	6.25	7.81	9.35	11.34	12.84	16.27	17.73
4	4.04	4.88	5.99	7.78	9.49	11.14	13.28	14.86	18.47	20.00
5	5.13	6.06	7.29	9.24	11.07	12.83	15.09	16.75	20.52	22.11
6	6.21	7.23	8.56	10.64	12.59	14.45	16.81	18.55	22.46	24.10
7	7.28	8.38	9.80	12.02	14.07	16.01	18.48	20.28	24.32	26.02
8	8.35	9.52	11.03	13.36	15.51	17.53	20.09	21.95	26.12	27.87
9	9.41	10.66	12.24	14.68	16.92	19.02	21.67	23.59	27.88	29.67
10	10.47	11.78	13.44	15.99	18.31	20.48	23.21	25.19	29.59	31.42
11	11.53	12.90	14.63	17.28	19.68	21.92	24.72	26.76	31.14	33.14
12	12.58	14.01	15.81	18.55	21.03	23.34	26.22	28.30	32.91	34.82
13	13.64	15.12	16.98	19.81	22.36	24.74	27.69	29.82	34.53	36.48
14	14.69	16.22	18.15	21.06	23.68	26.12	29.14	31.32	36.12	38.11
15	15.73	17.32	19.31	22.31	25.00	27.49	30.58	32.80	37.70	39.75
16	16.78	18.42	20.47	23.54	26.30	28.85	32.00	34.27	39.25	41.31
17	17.82	19.51	21.61	24.77	27.59	30.19	33.41	35.72	40.79	42.88
18	18.87	20.60	22.76	25.99	28.87	31.53	34.81	37.16	42.31	44.43
19	19.91	21.69	23.90	27.20	30.14	32.85	36.19	38.58	43.82	45.97
20	20.95	22.77	25.04	28.41	31.41	34.17	37.57	40.00	45.31	47.50
21	21.99	23.86	26.17	29.62	32.67	35.48	38.93	41.40	46.80	49.01
22	23.03	24.94	27.30	30.81	33.92	36.78	40.29	42.80	48.27	50.51
23	24.07	26.02	28.43	32.01	35.17	38.08	41.64	44.18	49.73	52.00
24	25.11	27.10	29.55	33.20	36.42	39.36	42.98	45.56	51.18	53.48
25	26.14	28.17	30.68	34.38	37.65	40.65	44.31	46.93	52.62	54.95
26	27.18	29.25	31.79	35.56	38.89	41.92	45.64	48.29	54.05	56.41
27	28.21	30.32	32.91	36.74	40.11	43.19	46.96	49.64	55.48	57.86
28	29.25	31.39	34.03	37.92	41.34	44.46	48.28	50.99	56.89	59.30
29	30.28	32.46	35.14	39.09	42.56	45.72	49.59	52.34	58.30	60.73
30	31.32	33.53	36.25	40.26	43.77	46.98	50.89	53.67	59.70	62.16
40	41.62	44.16	47.27	51.81	55.76	59.34	63.69	66.77	73.40	76.09
50	51.89	54.72	58.16	63.17	67.50	71.42	76.15	79.49	86.66	89.56
60	62.13	65.23	68.97	74.40	79.08	83.30	88.38	91.95	99.61	102.69
100	102.95	106.91	111.67	118.50	124.34	129.56	135.81	140.17	149.45	153.17

(9.4) Table for Binomial random variable $P(Bin(n, p) \leq k)$ if $p \leq 0.5$.
 If $p > 0.5$, then $P(Bin(n, p) \leq k) = P(Bin(n, 1-p) \geq n-k)$.

n	k	0.05	0.10	0.15	0.20	0.25	p	0.30	0.35	0.40	0.45	0.50
2	0	0.9025	0.8100	0.7225	0.6400	0.5625	0.4900	0.4225	0.3600	0.3025	0.2500	
3	0	0.9575	0.9900	0.9775	0.9600	0.9375	0.9100	0.8775	0.8400	0.7975	0.7500	
3	1	0.8574	0.7290	0.6141	0.5120	0.4219	0.3430	0.2746	0.2160	0.1664	0.1250	
2	1	0.9928	0.9720	0.9392	0.8960	0.8438	0.7840	0.7183	0.6480	0.5747	0.5000	
2	2	0.9999	0.9990	0.9966	0.9920	0.9844	0.9730	0.9571	0.9360	0.9089	0.8750	
4	0	0.8145	0.6561	0.5220	0.4096	0.3164	0.2401	0.1785	0.1296	0.0915	0.0625	
1	0.9860	0.9477	0.8905	0.8192	0.7383	0.6517	0.5630	0.4752	0.3910	0.3125		
2	0.9995	0.9963	0.9880	0.9728	0.9492	0.9163	0.8735	0.8208	0.7585	0.6875		
3	0	0.0000	0.9999	0.9995	0.9984	0.9961	0.9919	0.9850	0.9744	0.9590	0.9375	
5	0	0.7738	0.5905	0.4437	0.3277	0.2373	0.1681	0.1160	0.0778	0.0503	0.0313	
1	0.9774	0.9185	0.8352	0.7373	0.6328	0.5282	0.4284	0.3370	0.2562	0.1875		
2	0.9988	0.9914	0.9734	0.9421	0.8965	0.8369	0.7648	0.6826	0.5931	0.5000		
3	0	0.0000	0.9995	0.9978	0.9933	0.9844	0.9692	0.9460	0.8688	0.8125		
4	0	0.0000	1.0000	0.9999	0.9997	0.9990	0.9947	0.9898	0.9815	0.9781		
6	0	0.7351	0.5314	0.3771	0.2621	0.1780	0.1176	0.0754	0.0467	0.0277	0.0155	
1	0.9672	0.8857	0.7765	0.6554	0.5339	0.4202	0.3191	0.2333	0.1636	0.1094		
2	0.9978	0.9842	0.9527	0.9011	0.8306	0.7443	0.6471	0.5443	0.4415	0.3438		
3	0	0.9999	0.9987	0.9941	0.9830	0.9624	0.9295	0.8826	0.8208	0.7447	0.6563	
4	0	0.0000	0.9999	0.9996	0.9984	0.9954	0.9891	0.9777	0.9590	0.9308	0.8906	
5	0	1.0000	1.0000	0.9999	0.9998	0.9993	0.9982	0.9959	0.9917	0.9844		
7	0	0.6983	0.4783	0.3206	0.2097	0.1335	0.0824	0.0490	0.0280	0.0152	0.0078	
1	0.9556	0.8503	0.7166	0.5767	0.4449	0.3294	0.2338	0.1586	0.1024	0.0625		
2	0.9962	0.9743	0.9262	0.8520	0.7564	0.6471	0.5323	0.4199	0.3164	0.2266		
3	0	0.9998	0.9973	0.9878	0.9667	0.9294	0.8740	0.8002	0.7102	0.6083	0.5000	
4	0	1.0000	0.9998	0.9953	0.9871	0.9712	0.9444	0.9037	0.8471	0.7734		
5	0	1.0000	1.0000	0.9999	0.9996	0.9987	0.9962	0.9910	0.9812	0.9643	0.9375	
6	0	1.0000	1.0000	1.0000	0.9999	0.9998	0.9994	0.9984	0.9963	0.9922		
8	0	0.6634	0.4305	0.2725	0.1678	0.1001	0.0576	0.0319	0.0168	0.0084	0.0039	
1	0.9428	0.8131	0.6572	0.5033	0.3671	0.2553	0.1691	0.1064	0.0632	0.0352		
2	0.9942	0.9619	0.8948	0.7969	0.6785	0.5518	0.4278	0.3154	0.2201	0.1445		
3	0.9996	0.9950	0.9786	0.8862	0.8059	0.7064	0.5941	0.4770	0.3633			
4	0	1.0000	0.9996	0.9971	0.9896	0.9727	0.9420	0.8939	0.8263	0.7396	0.6367	
5	0	1.0000	1.0000	0.9998	0.9988	0.9887	0.9747	0.9502	0.9115	0.8555		
6	0	1.0000	1.0000	0.9999	0.9996	0.9987	0.9964	0.9915	0.9819	0.9643		
7	0	1.0000	1.0000	1.0000	0.9999	0.9998	0.9993	0.9983	0.9961			
9	0	0.6302	0.3874	0.2316	0.1342	0.0751	0.0404	0.0207	0.0101	0.0046	0.0020	
1	0.9288	0.7748	0.5995	0.4336	0.3003	0.1960	0.1211	0.0705	0.0385	0.0195		
2	0.9916	0.9470	0.8591	0.7382	0.6007	0.4628	0.3373	0.2318	0.1495	0.0898		
3	0.9994	0.9917	0.9661	0.9144	0.8343	0.7297	0.6089	0.4826	0.3614	0.2539		
4	0	1.0000	0.9991	0.9944	0.9804	0.9510	0.9012	0.8283	0.7334	0.6214	0.5000	
5	0	1.0000	0.9999	0.9994	0.9969	0.9900	0.9747	0.9464	0.9006	0.8342	0.7461	
6	0	1.0000	1.0000	0.9997	0.9987	0.9957	0.9888	0.9750	0.9502	0.9102		
7	0	1.0000	1.0000	0.9999	0.9996	0.9986	0.9962	0.9909	0.9805			
8	0	1.0000	1.0000	1.0000	0.9999	0.9999	0.9997	0.9992	0.9980			

n	k	0.05	0.10	0.15	0.20	0.25	p	0.30	0.35	0.40	0.45	0.50
10	0	0.5987	0.3487	0.1969	0.1074	0.0563	0.0282	0.0135	0.0060	0.0025	0.0010	
1	0	0.9139	0.7361	0.5443	0.3758	0.2440	0.1493	0.0860	0.0464	0.0233	0.0107	
2	0	0.9885	0.9298	0.8202	0.6778	0.5256	0.3828	0.2616	0.1673	0.0996	0.0547	
3	0	0.9990	0.9872	0.9500	0.8791	0.7759	0.6499	0.5138	0.3823	0.2660	0.1719	
4	0	0.9999	0.9984	0.9901	0.9672	0.9219	0.8497	0.7515	0.6331	0.5044	0.3770	
5	0	1.0000	0.9999	0.9986	0.9936	0.9803	0.9527	0.9051	0.8338	0.7584	0.6230	
6	0	1.0000	1.0000	0.9999	0.9991	0.9965	0.9894	0.9740	0.9452	0.8980	0.8281	
7	0	1.0000	1.0000	1.0000	0.9999	0.9996	0.9984	0.9952	0.9877	0.9726	0.9453	
8	0	1.0000	1.0000	1.0000	0.9999	0.9999	0.9995	0.9983	0.9955	0.9893		
9	0	1.0000	1.0000	1.0000	0.9999	0.9999	0.9997	0.9997	0.9998	0.9995		
10	0	1.0000	1.0000	1.0000	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999		
11	0	1.0000	1.0000	1.0000	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999		
12	0	1.0000	1.0000	1.0000	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999		
13	0	1.0000	1.0000	1.0000	0.9999	0.9999	0.9999	0.9999	0.9999	0.9999		

Table for Binomial random variable $P(Bin(n, p) \leq k)$ if $p \leq 0.5$.
If $p > 0.5$, then $P(Bin(n, p) \leq k) = P(Bin(n, 1-p) \geq n-k)$.

n	k	0.05	0.10	0.15	0.20	0.25	p	0.30	0.35	0.40	0.45	0.50
14	0	0.4877	0.2288	0.1028	0.0440	0.0178	0.0068	0.0024	0.0008	0.0002	0.0001	0.0000
1	0.8470	0.5284	0.3567	0.1979	0.1010	0.0475	0.0205	0.0081	0.0029	0.0009	0.0000	0.0000
2	0.9699	0.8416	0.6479	0.4481	0.2811	0.1608	0.0839	0.0398	0.0170	0.0065	0.0005	0.0001
3	0.9558	0.9559	0.8535	0.6982	0.5213	0.3552	0.2205	0.1243	0.0632	0.0287	0.0123	0.0012
4	0.9996	0.9908	0.9533	0.8702	0.7415	0.5842	0.4227	0.2793	0.1672	0.0898	0.0464	0.0184
5	1.0000	0.9985	0.9885	0.9561	0.8883	0.7805	0.6405	0.4859	0.3373	0.2120	0.1260	0.0596
6	1.0000	0.9998	0.9978	0.9884	0.9617	0.9067	0.8164	0.6925	0.5461	0.3953	0.2348	0.1071
7	1.0000	1.0000	0.9997	0.9976	0.9897	0.9247	0.8499	0.7414	0.6047	0.4478	0.2902	0.1662
8	1.0000	1.0000	1.0000	0.9996	0.9978	0.9917	0.9757	0.8811	0.7830	0.6405	0.4743	0.3145
9	1.0000	1.0000	1.0000	1.0000	0.9997	0.9983	0.9940	0.9825	0.9574	0.9006	0.8011	0.6626
10	1.0000	1.0000	1.0000	1.0000	1.0000	0.9998	0.9989	0.9961	0.9886	0.9713	0.9617	0.9081
11	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9994	0.9978	0.9935	0.9911	0.9755
12	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9997	0.9999	0.9970	0.9894
13	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
14	0	0.4633	0.2059	0.0874	0.0352	0.0134	0.0047	0.0016	0.0005	0.0001	0.0000	0.0000
1	0.8290	0.5490	0.3186	0.1671	0.0802	0.0353	0.0142	0.0052	0.0017	0.0005	0.0000	0.0000
2	0.9638	0.8159	0.6042	0.3980	0.2361	0.1268	0.0617	0.0271	0.0107	0.0037	0.0000	0.0000
3	0.9445	0.9444	0.8227	0.6482	0.4613	0.2969	0.1727	0.0905	0.0424	0.0176	0.0000	0.0000
4	0.9994	0.9873	0.9383	0.8358	0.6865	0.5155	0.3519	0.2173	0.1204	0.0592	0.0000	0.0000
5	0.9999	0.9978	0.9832	0.9389	0.8516	0.7216	0.5643	0.4032	0.2608	0.1509	0.0000	0.0000
6	1.0000	0.9997	0.9964	0.9819	0.9434	0.8689	0.7548	0.6098	0.4522	0.3036	0.2000	0.1000
7	1.0000	1.0000	0.9994	0.9958	0.9827	0.9500	0.8868	0.7869	0.6535	0.5000	0.3838	0.2000
8	1.0000	1.0000	0.9999	0.9992	0.9958	0.9848	0.9578	0.9050	0.8182	0.6964	0.5344	0.3000
9	1.0000	1.0000	1.0000	0.9990	0.9992	0.9963	0.9876	0.9662	0.9231	0.8491	0.7210	0.4000
10	1.0000	1.0000	1.0000	0.9999	0.9993	0.9972	0.9907	0.9745	0.9408	0.8743	0.7597	0.4743
11	1.0000	1.0000	1.0000	1.0000	0.9999	0.9995	0.9981	0.9937	0.9824	0.8894	0.7936	0.4743
12	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
13	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
14	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
15	0	0.4401	0.1853	0.0743	0.0281	0.0100	0.0033	0.0010	0.0003	0.0001	0.0000	0.0000
1	0.8108	0.5147	0.2839	0.1407	0.0635	0.0261	0.0098	0.0033	0.0010	0.0003	0.0000	0.0000
2	0.9571	0.7892	0.5614	0.3518	0.1971	0.0994	0.0451	0.0183	0.0066	0.0021	0.0000	0.0000
3	0.9380	0.9316	0.7899	0.5981	0.4050	0.2459	0.1339	0.0651	0.0281	0.0106	0.0000	0.0000
4	0.9991	0.9830	0.9209	0.7982	0.6302	0.4499	0.2892	0.1666	0.0853	0.0384	0.0000	0.0000
5	0.9999	0.9967	0.9765	0.9183	0.8103	0.6598	0.4900	0.3288	0.1976	0.1051	0.0000	0.0000
6	1.0000	0.9995	0.9944	0.9733	0.9204	0.8247	0.6881	0.5272	0.3660	0.2272	0.1000	0.0000
7	1.0000	0.9999	0.9930	0.9729	0.9256	0.8406	0.7161	0.5629	0.4018	0.2704	0.1443	0.0704
8	1.0000	1.0000	0.9998	0.9985	0.9743	0.9329	0.8579	0.7441	0.5982	0.4551	0.2631	0.1332
9	1.0000	1.0000	1.0000	0.9998	0.9929	0.9771	0.9417	0.8759	0.7728	0.6980	0.5648	0.3854
10	1.0000	1.0000	1.0000	1.0000	0.9997	0.9938	0.9809	0.9514	0.8949	0.8029	0.6768	0.4739
11	1.0000	1.0000	1.0000	1.0000	0.9997	0.9987	0.9951	0.9851	0.9616	0.8983	0.7667	0.5778
12	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
13	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
14	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
15	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999

Table for Binomial random variable $P(Bin(n, p) \leq k)$ if $p \leq 0.5$.
If $p > 0.5$, then $P(Bin(n, p) \leq k) = P(Bin(n, 1-p) \geq n-k)$.

n	k	0.05	0.10	0.15	0.20	0.25	p	0.30	0.35	0.40	0.45	0.50
17	0	0.4181	0.1668	0.0631	0.0225	0.0075	0.0023	0.0007	0.0002	0.0000	0.0000	0.0000
1	0.7922	0.4818	0.2525	0.1182	0.0501	0.0193	0.0067	0.0021	0.0006	0.0000	0.0000	0.0000
2	0.9497	0.7618	0.5198	0.3096	0.1637	0.0774	0.0327	0.0123	0.0041	0.0012	0.0000	0.0000
3	0.9912	0.9174	0.7556	0.5489	0.3530	0.2019	0.1028	0.0464	0.0184	0.0064	0.0000	0.0000
4	0.9988	0.9779	0.9013	0.7582	0.5739	0.3887	0.2348	0.1260	0.0596	0.0245	0.0000	0.0000
5	0.9999	0.9953	0.9681	0.8943	0.7653	0.5968	0.4197	0.2639	0.1471	0.0717	0.0000	0.0000
6	1.0000	0.9992	0.9623	0.8929	0.7752	0.6188	0.4478	0.2902	0.1662	0.0833	0.0000	0.0000
7	1.0000	0.9999	0.9883	0.8954	0.8585	0.7872	0.6405	0.4743	0.3145	0.1923	0.0000	0.0000
8	1.0000	1.0000	1.0000	0.9997	0.9970	0.9970	0.9970	0.9970	0.9970	0.9970	0.9970	0.9970
9	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
10	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
11	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
12	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
13	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
14	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
15	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
16	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000

(9.5) Table for Poisson random variable $P(Po(\mu) \leq k)$.

k	μ									
k	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	1.0
0	0.9048	0.8187	0.7408	0.6703	0.6065	0.5488	0.4966	0.4493	0.4066	0.3679
1	0.9953	0.9825	0.9631	0.9384	0.9098	0.8781	0.8442	0.8088	0.7725	0.7358
2	0.9998	0.9989	0.9964	0.9921	0.9856	0.9769	0.9659	0.9526	0.9371	0.9197
3	1.0000	0.9999	0.9997	0.9992	0.9982	0.9966	0.9942	0.9909	0.9865	0.9810
4	1.0000	1.0000	0.9999	0.9998	0.9996	0.9992	0.9986	0.9977	0.9963	0.9956
5	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9998	0.9994
6	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
7	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
k	1.1	1.2	1.3	1.4	1.5	1.6	1.7	1.8	1.9	2.0
0	0.9329	0.3012	0.2725	0.2466	0.2231	0.2019	0.1827	0.1653	0.1496	0.1353
1	0.6890	0.6626	0.6268	0.5918	0.5578	0.5249	0.4932	0.4628	0.4337	0.4060
2	0.9004	0.8795	0.8571	0.8335	0.8088	0.7834	0.7572	0.7306	0.7037	0.6767
3	0.9743	0.9662	0.9569	0.9463	0.9344	0.9212	0.9068	0.8913	0.8747	0.8571
4	0.9946	0.9923	0.9893	0.9857	0.9814	0.9763	0.9704	0.9636	0.9559	0.9473
5	0.9990	0.9985	0.9978	0.9968	0.9955	0.9940	0.9920	0.9896	0.9834	0.9775
6	0.9999	0.9997	0.9996	0.9994	0.9991	0.9987	0.9981	0.9974	0.9966	0.9955
7	1.0000	1.0000	0.9999	0.9999	0.9998	0.9997	0.9996	0.9994	0.9992	0.9989
8	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
9	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
k	3.2	3.4	3.6	3.8	4.0	4.2	4.4	4.6	4.8	5.0
0	0.0408	0.0334	0.0273	0.0224	0.0183	0.0150	0.0123	0.0101	0.0082	0.0067
1	0.1712	0.1468	0.1257	0.1074	0.0916	0.0780	0.0663	0.0563	0.0477	0.0404
2	0.3799	0.3397	0.3027	0.2689	0.2381	0.2102	0.1851	0.1626	0.1425	0.1247
3	0.6025	0.5584	0.5152	0.4735	0.4335	0.3954	0.3594	0.3257	0.2942	0.2650
4	0.7806	0.7442	0.7064	0.6678	0.6288	0.5898	0.5512	0.5132	0.4763	0.4405
5	0.8946	0.8705	0.8441	0.8156	0.7851	0.7531	0.7199	0.6858	0.6510	0.6160
6	0.9554	0.9421	0.9267	0.9091	0.8893	0.8675	0.8436	0.8180	0.7908	0.7622
7	0.9832	0.9769	0.9692	0.9559	0.9489	0.9361	0.9214	0.9049	0.8867	0.8666
8	0.9943	0.9917	0.9883	0.9840	0.9786	0.9721	0.9642	0.9549	0.9442	0.9319
9	0.9995	0.9992	0.9987	0.9981	0.9972	0.9959	0.9943	0.9922	0.9896	0.9863
k	5.2	5.4	5.6	5.8	6.0	6.5	7.0	7.5	8.0	8.5
0	0.0055	0.0045	0.0037	0.0030	0.0025	0.0015	0.0009	0.0006	0.0003	0.0002
1	0.0342	0.0289	0.0244	0.0206	0.0174	0.0113	0.0073	0.0047	0.0030	0.0019
2	0.1088	0.0948	0.0824	0.0715	0.0620	0.0430	0.0296	0.0203	0.0138	0.0093
3	0.2381	0.2133	0.1906	0.1700	0.1512	0.1118	0.0818	0.0591	0.0424	0.0301
4	0.4061	0.3733	0.3422	0.3127	0.2851	0.2237	0.1730	0.1321	0.0966	0.0744
5	0.5809	0.5461	0.5119	0.4783	0.4457	0.3690	0.3007	0.2414	0.1912	0.1496
6	0.7324	0.7017	0.6703	0.6384	0.6063	0.5265	0.4497	0.3782	0.3134	0.2562
7	0.8449	0.8149	0.7970	0.7710	0.7440	0.6728	0.5987	0.5246	0.4530	0.3856
8	0.9181	0.9027	0.8857	0.8672	0.8472	0.7916	0.7291	0.6620	0.5925	0.5231
9	0.9603	0.9512	0.9409	0.9292	0.9161	0.8774	0.8305	0.7764	0.7166	0.6530
10	0.9823	0.9775	0.9718	0.9651	0.9574	0.9332	0.9015	0.8622	0.8159	0.7634
11	0.9904	0.9875	0.9841	0.9799	0.9661	0.9467	0.9208	0.8881	0.8487	0.8047
12	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
k	2.1	2.2	2.3	2.4	2.5	2.6	2.7	2.8	2.9	3.0
0	0.1225	0.1108	0.1003	0.0907	0.0821	0.0743	0.0672	0.0608	0.0550	0.0498
1	0.3796	0.3546	0.3309	0.3084	0.2873	0.2674	0.2487	0.2311	0.2146	0.1991
2	0.6496	0.6227	0.5960	0.5597	0.5348	0.5184	0.4936	0.4695	0.4460	0.4232
3	0.8386	0.8194	0.7787	0.7576	0.7360	0.7141	0.6919	0.6696	0.6474	0.6141
4	0.9379	0.9275	0.9162	0.9041	0.8912	0.8774	0.8477	0.8318	0.8153	0.7894
5	0.9796	0.9751	0.9700	0.9643	0.9580	0.9510	0.9433	0.9349	0.9258	0.9161
6	0.9941	0.9925	0.9884	0.9828	0.9794	0.9756	0.9713	0.9665	0.9623	0.9573
7	0.9985	0.9974	0.9967	0.9958	0.9947	0.9934	0.9919	0.9901	0.9881	0.9851
8	0.9997	0.9995	0.9994	0.9989	0.9985	0.9981	0.9976	0.9969	0.9962	0.9954
9	0.9999	0.9999	0.9999	0.9998	0.9997	0.9996	0.9995	0.9994	0.9993	0.9992
10	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
11	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
12	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000

Table for Poisson random variable $P(Po(\mu) \leq k)$.

Table for Poisson random variable $P(Po(\mu) \leq k)$.

k	9.0	9.5	10.0	11.0	12.0	13.0	14.0	15.0	16.0	17.0	μ
0	0.0001	0.0001	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
1	0.0012	0.0008	0.0005	0.0002	0.0001	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
2	0.0062	0.0042	0.0028	0.0012	0.0005	0.0002	0.0001	0.0000	0.0000	0.0000	0.0000
3	0.0212	0.0149	0.0103	0.0049	0.0023	0.0011	0.0005	0.0002	0.0001	0.0000	0.0000
4	0.0550	0.0403	0.0293	0.0151	0.0076	0.0037	0.0018	0.0009	0.0004	0.0002	0.0000
5	0.1157	0.0885	0.0671	0.0375	0.0203	0.0107	0.0055	0.0028	0.0014	0.0007	0.0000
6	0.2068	0.1649	0.1301	0.0786	0.0458	0.0259	0.0142	0.0076	0.0040	0.0021	0.0000
7	0.3239	0.2687	0.2202	0.1432	0.0895	0.0540	0.0316	0.0180	0.0100	0.0054	0.0000
8	0.4557	0.3918	0.3328	0.2320	0.1550	0.0998	0.0621	0.0374	0.0220	0.0126	0.0000
9	0.5874	0.5218	0.4579	0.3405	0.2424	0.1658	0.1094	0.0699	0.0433	0.0261	0.0000
10	0.7060	0.6453	0.5830	0.4599	0.3472	0.2517	0.1757	0.1185	0.0774	0.0491	0.0000
11	0.8030	0.7520	0.6968	0.5793	0.4616	0.3532	0.2600	0.1848	0.1270	0.0847	0.0000
12	0.8758	0.8364	0.7916	0.6887	0.5760	0.4631	0.3585	0.2676	0.1931	0.1350	0.0000
13	0.9261	0.8981	0.8645	0.7813	0.6815	0.5730	0.4644	0.3632	0.2745	0.2009	0.0000
14	0.9585	0.9400	0.9165	0.8540	0.7720	0.6751	0.5704	0.4657	0.3675	0.2808	0.0000
15	0.9780	0.9665	0.9513	0.9074	0.8444	0.7636	0.6694	0.5681	0.4667	0.3715	0.0000
16	0.9889	0.9823	0.9730	0.9441	0.8987	0.8355	0.7559	0.6641	0.5660	0.4677	0.0000
17	0.9947	0.9911	0.9857	0.9678	0.9370	0.8905	0.8272	0.7489	0.6593	0.5640	0.0000
18	0.9976	0.9957	0.9928	0.9823	0.9626	0.9302	0.8826	0.8195	0.7423	0.6550	0.0000
19	0.9989	0.9980	0.9965	0.9907	0.9787	0.9573	0.9235	0.8752	0.8122	0.7363	0.0000
20	0.9996	0.9991	0.9984	0.9953	0.9884	0.9750	0.9521	0.9170	0.8682	0.8055	0.0000
21	0.9998	0.9996	0.9993	0.9977	0.9939	0.9859	0.9712	0.9469	0.9108	0.8615	0.0000
22	0.9999	0.9999	0.9997	0.9990	0.9970	0.9924	0.9833	0.9673	0.9418	0.9047	0.0000
23	1.0000	0.9999	0.9999	0.9995	0.9985	0.9960	0.9907	0.9805	0.9633	0.9367	0.0000
24	1.0000	1.0000	1.0000	0.9998	0.9993	0.9980	0.9950	0.9888	0.9777	0.9594	0.0000
25	1.0000	1.0000	1.0000	0.9999	0.9997	0.9974	0.9938	0.9869	0.9748	0.9548	0.0000
26	1.0000	1.0000	1.0000	1.0000	0.9999	0.9995	0.9987	0.9967	0.9925	0.9848	0.0000
27	1.0000	1.0000	1.0000	1.0000	0.9999	0.9998	0.9994	0.9983	0.9959	0.9912	0.0000
28	1.0000	1.0000	1.0000	1.0000	0.9999	0.9997	0.9991	0.9978	0.9950	0.9900	0.0000
29	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9996	0.9989	0.9973	0.9947	0.0000
30	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9998	0.9994	0.9986	0.9976	0.0000
31	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9997	0.9993	0.9980	0.0000
32	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9996	0.0000
33	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9998	0.0000
34	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.0000
35	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.0000